

Conditional Distributions and Densities

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Outline

- Conditional Expected Value
- Chapman – Kolmogorov Equation
- Sample Mean and Variance
- Estimating Mean and Variance of Random Data
- Alternative Definition For Probability Density Function

Conditional Distribution of Y given event m

$$F_Y(y | m) = P\{Y \leq y | m\} = \frac{P\{Y \leq y \cap m\}}{P(m)}$$

For $m = \{X \leq x\}$,

$$F_Y(y | X \leq x) = \frac{P\{X \leq x \cap Y \leq y\}}{P\{X \leq x\}} = \frac{F_{XY}(x, y)}{F_X(x)}$$

$$f_Y(y | X \leq x) = \frac{\partial F_{XY}(x, y)}{\partial y} = \frac{1}{F_X(x)} \int_{-\infty}^x f_{XY}(x_1, y) dx_1$$

Similarly:

$$F_Y(y | x_1 < X \leq x_2) = \frac{\int_{x_1}^{x_2} f_{XY}(x, y) dx}{F_X(x_2) - F_X(x_1)}$$

Conditional Distribution & Density given X = x

Noting $F_Y(y | X = x) = \lim_{\Delta x \rightarrow 0} F_Y(y | x < X \leq x + \Delta x)$,

$$F_Y(y | x = x) = \lim_{\Delta x \rightarrow 0} \frac{F_{XY}(x + \Delta x, y) - F_{XY}(x, y)}{F_X(x + \Delta x) - F_X(x)} = \frac{\partial F_{XY}(x, y)}{dF_X(x)}$$

$$F_Y(y | x = x) = \frac{\int_{-\infty}^y f_{XY}(x, y_1) dy_1}{f_X(x)}$$

$$f_Y(y | x) = f_Y(y | x = x) = \frac{f_{XY}(x, y)}{f_X(x)}$$

Similarly:

$$F_X(x | y) = \frac{\int_{-\infty}^x f_{XY}(x_1, y) dx_1}{f_Y(y)}$$

$$f_X(x | y) = \frac{f_{XY}(x, y)}{f_Y(y)}$$

Conditional Statistics Clarkson University

Conditional expected value of a function of a random variable g(Y)

$$E\{g(Y)|m\} = \int_{-\infty}^{+\infty} g(y)f_Y(y|m)dy$$

$$E\{g(Y)|X=x\} = \int_{-\infty}^{+\infty} g(y)f_Y(y|x)dy$$

Conditional Expected Value given X = x

$$E\{g(Y)|X=x\} = \frac{1}{f_X(x)} \int_{-\infty}^{+\infty} g(y)f_{XY}(x,y)dy$$

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Chapman - Kolmogorov Equation Clarkson University

Noting $f_X(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$,

$$f_X(x|y,z) = \frac{f_{XY}(x,y|z)}{f_Y(y|z)}$$

or

$$f_{XY}(x,y|z) = f_X(x|y,z)f_Y(y|z)$$

Integrating over y:

$$f_X(x|z) = \int_{-\infty}^{+\infty} f_X(x|y,z)f_Y(y|z)dy$$

For Markov processes:

$$f_X(x|y,z) = f_X(x|y)$$

$$f_X(x|z) = \int_{-\infty}^{+\infty} f_X(x|y)f_Y(y|z)dy$$

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Sample Mean & Variance Clarkson University

$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\bar{V} = \frac{(X_1 - \bar{X})^2 + \dots + (X_n - \bar{X})^2}{n}$$

If X_i have the same mean & variance and form a sequence of uncorrelated random variables:

$$E\{\bar{X}\} = \eta$$

$$\sigma_{\bar{X}}^2 = \frac{\sigma^2}{n}$$

$$E\{\bar{V}\} = \frac{n-1}{n} \sigma^2$$

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Statistics of Sample Mean & Variance Clarkson University

When X_i are jointly normal with

$$f(x_1, \dots, x_n) = \frac{1}{(2\pi)^{\frac{n}{2}} \sigma^n} \exp\left\{-\frac{x_1^2 + x_2^2 + \dots + x_n^2}{2\sigma^2}\right\}, \text{ density functions}$$

of $\bar{X} = \frac{1}{n} \sum_{j=1}^n X_j$ and $\bar{V} = \frac{1}{n} \sum_{j=1}^n (X_j - \bar{X})^2$ become

$$f_{\bar{X}}(\bar{x}) = \frac{1}{\sqrt{2\pi\sigma^2/n}} \exp\left\{-\frac{n\bar{x}^2}{2\sigma^2}\right\}$$

$$f_{\bar{V}}(v) = \frac{1}{2^{\frac{(n-1)}{2}} \left(\frac{\sigma}{\sqrt{n}}\right)^{n-1} \Gamma\left(\frac{n-1}{2}\right)} v^{\frac{(n-3)}{2}} e^{-\frac{nv}{2\sigma^2}} U(v)$$

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Chi- and Chi - Square Statistics Clarkson University

$$\chi = \sqrt{\sum_{j=1}^n X_j^2}$$

$$\chi^2 = Y = \sum_{j=1}^n X_j^2$$

Density functions of χ and $\chi^2 = Y$:

$$f_{\chi}(\chi) = \frac{2}{2^{\frac{n}{2}} \sigma^n \Gamma\left(\frac{n}{2}\right)} \chi^{n-1} e^{-\frac{\chi^2}{2\sigma^2}} U(\chi)$$

$$f_{\chi^2}(y) = \frac{1}{2^{\frac{n}{2}} \sigma^n \Gamma\left(\frac{n}{2}\right)} y^{\frac{(n-2)}{2}} e^{-\frac{y}{2\sigma^2}} U(y)$$

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Estimating Mean & Variance of Random Data Clarkson University

$$\bar{X} = \frac{\sum X_i}{n}$$

$$S^2 = \frac{\sum (X_i - \bar{X})^2}{n-1}$$

If X_i have η and σ as mean and variance:

$$E\{\bar{X}\} = \eta$$

$$\sigma_{\bar{X}}^2 = \frac{\sigma^2}{n}$$

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Central Limit Theorem Clarkson University

Theorem: If \bar{X} is the mean of a random sample of size n taken from a population having mean η and variance σ^2 , then $Z = \frac{\bar{X} - \eta}{\sigma/\sqrt{n}}$ is a random variable whose distribution approaches a standard normal distribution as $n \rightarrow \infty$, i.e.

$$f_z(z) = \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{z^2}{2}\right\}$$

$$P\{|Z| \leq z\} = 2\text{erf}(z)$$

Note:

$$P\{|Z| \leq 1\} \approx 0.68$$

$$P\{|Z| \leq 2\} = 0.85$$

$$P\{|Z| \leq 3\} = 0.997$$

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Sample Size Clarkson University

Size of a Sample for a Required Accuracy

Let error $E = |\bar{X} - \eta|$ and the set $|Z| \leq z$ lead to $E \leq z \frac{\sigma}{\sqrt{n}}$. The sample size needed is $n = \frac{z^2 \sigma^2}{E^2}$; if this is the sample size, then with probability of $2\text{erf}(z)$ the error will not be more than E .

Example: Let $z = 3$, $\sigma = 2$, $E = 0.01$, then $n = (9)(4)/10^{-4}$ data points are needed to estimate mean with probability 0.997 and error < 0.01

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Alternative Probability Density Function

Stratonovich definition of Probability Density Function



$$f_X(x) = E\{\delta(X - x)\}$$

$$E\{\delta(X - x)\} = \int_{-\infty}^{+\infty} \delta(x_1 - x) f_X(x_1) dx_1 = f_X(x)$$

$$E\{g(x)\} = E\left\{\int_{-\infty}^{+\infty} g(x) \delta(x - X) f_X(x) dx\right\} \\ = \int_{-\infty}^{+\infty} g(x) E\{\delta(x - X)\} dx \\ = \int_{-\infty}^{+\infty} g(x) f_X(x) dx$$

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Alternative Probability Density Function

Transformation of Random Variable Using the New Definition, $Y = g(X)$

$$f_Y(y) = \int_{-\infty}^{+\infty} \delta[g(x) - y] f_X(x) dx$$

$$\delta[g(x) - y] = \sum_j \frac{\delta(x - x_j)}{|g'(x_j)|}$$



$$f_Y(y) = \int_{-\infty}^{+\infty} \sum_j \delta(x - x_j) f_X(x) dx = \sum_j \frac{f_X(x_j)}{|g'(x_j)|}$$

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Concluding Remarks

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Thank you!

Questions?

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